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قسم الاحصاء

التوزيعات الاحتمالية
Probability Distributions

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Chapter Three

Some Continuous Probability Distributions

1. Continuous Uniform Distribution

A random variable X is said to have a continuous uniform distribution on the interval (a, b) if its probability density function (p.d.f.) is given by:

$$f(x, a, b) = \begin{cases} \frac{1}{b-a} & a \leq x \leq b \\ 0 & o.w. \end{cases}$$

Where the parameters a, b satisfy $-\infty < a < b < \infty$, we will write $X \sim U(a, b)$.

1. $f(x, a, b) \geq 0$
2. $\int_{\forall x} f(x, a, b) dx = 1$

$$\int_a^b \frac{1}{b-a} dx = \frac{1}{b-a} \int_a^b dx = \frac{1}{b-a} [x]_a^b = \frac{b-a}{b-a} = 1$$

3. The **mean** of the r.v. $X \sim U(a, b)$ can be obtain as follows

$$\mu_x = E(x) = \frac{a+b}{2}$$

4. The **variance** of the r.v. $X \sim U(a, b)$ can be obtain as follows

$$var(x) = \sigma_x^2 = E(x^2) - (E(x))^2 = \frac{(b-a)^2}{12}$$

5. The Cumulative Distribution Function of X is given by:

$$\begin{aligned} F(x) = P(X \leq x) &= \int_{-\infty}^x f(t) dt = \int_a^x \frac{1}{b-a} dt = \frac{1}{b-a} \int_a^x dt \\ &= \frac{1}{b-a} [t]_a^x = \frac{x-a}{b-a} \end{aligned}$$

$$F(x) = \begin{cases} 0 & x < a \\ \frac{x-a}{b-a} & a \leq x < b \\ 1 & x \geq b \end{cases}$$

Ex.1: If $X \sim U(0,10)$.

1. Write the probability density function of X .
2. $P(x < 3)$, $P(x > 6)$, and $P(3 < x < 8)$ using the p.d.f.
3. What is the cumulative distribution function of X .
4. $P(x < 3)$, $P(x > 6)$, and $P(3 < x < 8)$ using the c.d.f.
5. Find the mean and the variance of X .

Solution:

$$1. f(x) = \begin{cases} \frac{1}{10} & 0 \leq x \leq 10 \\ 0 & \text{o.w.} \end{cases}$$

$$2. P(x < 3) = \int_0^3 f(x) dx = \int_0^3 \frac{1}{10} dx = \frac{1}{10} [x]_0^3 = \frac{3}{10}$$

$$P(x > 6) = \int_6^{10} f(x) dx = \int_6^{10} \frac{1}{10} dx = \frac{1}{10} [x]_6^{10} = \frac{10-6}{10} = \frac{4}{10}$$

$$P(3 < x < 8) = \int_3^8 f(x) dx = \int_3^8 \frac{1}{10} dx = \frac{1}{10} [x]_3^8 = \frac{8-3}{10} = \frac{5}{10}$$

$$3. F(x) = \begin{cases} 0 & x < a \\ \frac{x-a}{b-a} & a \leq x < b \\ 1 & x \geq b \end{cases}$$

$$F(x) = \begin{cases} 0 & x < 0 \\ \frac{x}{10} & 0 \leq x < 10 \\ 1 & x \geq 10 \end{cases}$$

4. We can use $F(x)$ to compute the probabilities above.

$$P(x < 3) = F(3) - F(0) = \frac{3}{10} - 0 = \frac{3}{10}$$

$$P(x > 6) = F(10) - F(6) = 1 - \frac{6}{10} = \frac{4}{10}$$

$$P(3 < x < 8) = F(8) - F(3) = \frac{8}{10} - \frac{3}{10} = \frac{5}{10}$$

$$5. \mu_x = \frac{a+b}{2} = \frac{0+10}{2} = 5$$

$$\text{var}(x) = \sigma_x^2 = \frac{(b-a)^2}{12} = \frac{(10-0)^2}{12} = \frac{100}{12}$$

Ex.2: A bus arrives at a station at a random time between 8:10 AM and 8:30 AM. Assume that the arrival time is uniformly distributed over this interval.

1. Write the probability density function of X .
2. Find the probability that the bus arrives before 8:15 AM.
3. Find the probability that the bus arrives after 8:20 AM.
4. Find $P(13 < x < 18)$.
5. What is the cumulative distribution function of X .
6. Find the mean and the variance of X .

Solution:

$$1. X \sim U(10,30)$$

$$f(x) = \begin{cases} \frac{1}{20} & 10 \leq x \leq 30 \\ 0 & \text{o.w.} \end{cases}$$

$$2. P(x < 15) = \int_{10}^{15} f(x) dx = \int_{10}^{15} \frac{1}{20} dx = \frac{1}{20} [x]_{10}^{15} = \frac{15-10}{20} = \frac{5}{20}$$

$$3. P(x > 20) = \int_{20}^{30} f(x) dx = \int_{20}^{30} \frac{1}{20} dx = \frac{1}{20} [x]_{20}^{30} = \frac{30-20}{20} = \frac{10}{20}$$

$$4. P(13 < x < 18) = \int_{13}^{18} f(x) dx = \int_{13}^{18} \frac{1}{20} dx = \frac{1}{20} [x]_{13}^{18} = \frac{18-13}{20} = \frac{5}{20}$$

$$5. F(x) = \begin{cases} 0 & x < 10 \\ \frac{x-10}{20} & 10 \leq x < 30 \\ 1 & x \geq 30 \end{cases}$$

$$6. \mu_x = \frac{a+b}{2} = \frac{10+30}{2} = \frac{40}{2} = 20$$

$$\text{var}(x) = \sigma_x^2 = \frac{(b-a)^2}{12} = \frac{(30-10)^2}{12} = \frac{400}{12}$$

2. The Normal Distribution

A random variable X is said to have a normal distribution with parameters μ and σ^2 if its probability density function (p.d.f.) is given by:

$$f(x, \mu, \sigma^2) = \begin{cases} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} & -\infty < x < \infty \\ 0 & \text{o.w.} \end{cases}$$

Where $-\infty < \mu < \infty$ and $0 < \sigma^2 < \infty$, the parameter μ represents the mean of x and the parameter σ^2 represents the variance of x . We write $X \sim N(\mu, \sigma^2)$.

1. $f(x, \mu, \sigma^2) \geq 0$
2. $\int_{\forall x} f(x, \mu, \sigma^2) dx = 1$

$$\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx = \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{\infty} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

$$\text{Let } z = \frac{x-\mu}{\sigma} \Rightarrow z\sigma = x - \mu$$

$$\sigma dz = dx$$

$$\frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{\infty} e^{-\frac{z^2}{2}} \sigma dz$$

$$\frac{\sigma}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{z^2}{2}} dz$$

$$\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{z^2}{2}} dz$$

$$\text{Recall that: } \int_{-\infty}^{\infty} e^{-\frac{z^2}{2}} dz = \sqrt{2\pi}$$

$$\frac{1}{\sqrt{2\pi}} \sqrt{2\pi} = 1$$

3. The **mean** of the r.v. $X \sim N(\mu, \sigma^2)$ can be obtain as follows

$$\mu_x = E(x) = \mu$$

4. The **variance** of the r.v. $X \sim N(\mu, \sigma^2)$ can be obtain as follows

$$\text{var}(x) = \sigma_x^2 = E(x^2) - (E(x))^2 = \sigma^2$$

5. The Cumulative Distribution Function of X is given by:

$$F(x) = P(X \leq x) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}\sigma^2} e^{-\frac{(t-\mu)^2}{2\sigma^2}} dt$$

Remark:

1. If $X \sim N(\mu, \sigma^2)$, then $z = \frac{x-\mu}{\sigma}$ is a **Standard Normal Distribution** with mean 0 and variance 1 $z \sim SN(0,1)$.
2. The probability density function (p.d.f.) of the Standard Normal Distribution is given by:

$$f(z) = \begin{cases} \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} & -\infty < z < \infty \\ 0 & \text{other wise} \end{cases}$$

3. $P(a < z < b) = \phi(b) - \phi(a)$. $\phi(b), \phi(a)$ are calculate from the tables.
4. $P(z \leq a) = \phi(a)$
5. $P(z > a) = 1 - P(z \leq a) = 1 - \phi(a)$
6. $P(z < -a) = P(z > a) = 1 - P(z \leq a) = 1 - \phi(a)$
7. $\phi(-a) = 1 - \phi(a)$

Ex.3: Let $X \sim N(4,25)$. $\phi(0.4) = 0.6554$, $\phi(0.2) = 0.5793$

1. Write the p.d.f. of X .
2. Find the mean and the variance of X .
3. Find $P(x < 6)$.
4. Find $P(3 < x < 5)$.
5. Find $P(x \leq 2)$.

Solution:

1. $X \sim N(4,25)$

$$f(x, \mu, \sigma^2) = \begin{cases} \frac{1}{\sqrt{2\pi \cdot 25}} e^{-\frac{(x-4)^2}{2(25)}} & -\infty < x < \infty \\ 0 & \text{o.w.} \end{cases}$$

$$f(x, 4, 25) = \begin{cases} \frac{1}{\sqrt{50\pi}} e^{-\frac{(x-4)^2}{50}} & -\infty < x < \infty \\ 0 & \text{o.w.} \end{cases}$$

2. $\mu_x = E(x) = \mu = 4$, $\text{var}(x) = \sigma_x^2 = \sigma^2 = 25$

$$\begin{aligned} 3. P(x < 6) &= P\left(\frac{x-\mu}{\sigma} < \frac{6-\mu}{\sigma}\right) = P\left(z < \frac{6-4}{5}\right) = P\left(z < \frac{2}{5}\right) \\ &= P(z < 0.4) = \phi(0.4) = 0.6554 \end{aligned}$$

$$\begin{aligned} 4. P(3 < x < 5) &= P\left(\frac{3-\mu}{\sigma} < \frac{x-\mu}{\sigma} < \frac{5-\mu}{\sigma}\right) = P\left(\frac{3-4}{5} < z < \frac{5-4}{5}\right) \\ &= P(-0.2 < z < 0.2) = \phi(0.2) - \phi(-0.2) \\ &= \phi(0.2) - [1 - \phi(0.2)] = 0.5793 - [1 - 0.5793] = 0.1586 \end{aligned}$$

$$\begin{aligned} 5. P(x \leq 2) &= P\left(\frac{x-\mu}{\sigma} \leq \frac{2-\mu}{\sigma}\right) = P\left(z < \frac{2-4}{5}\right) = P\left(z < \frac{-2}{5}\right) \\ &= P(z < -0.4) = P(z > 0.4) = 1 - P(z \leq 0.4) \\ &= 1 - \phi(0.4) = 1 - 0.6554 = 0.3446 \end{aligned}$$

Or

$$\begin{aligned} P(x \leq 2) &= P\left(\frac{x-\mu}{\sigma} \leq \frac{2-\mu}{\sigma}\right) = P\left(z < \frac{2-4}{5}\right) = P\left(z < \frac{-2}{5}\right) \\ &= P(z < -0.4) = \phi(-0.4) \\ &= 1 - \phi(0.4) = 1 - 0.6554 = 0.3446 \end{aligned}$$

Ex.4: Suppose the temperature in Iraq during July follows a normal distribution with mean 45 °C and standard deviation 2.5 °C. Find the probability that the temperature is:

1. Between 46.7 °C and 48.1 °C.
2. Greater than 47.5 °C.

Solution:

$$X \sim N(45, (2.5)^2) \Rightarrow X \sim N(45, 6.25)$$

$$f(x, \mu, \sigma^2) = \begin{cases} \frac{1}{\sqrt{12.5 \pi}} e^{-\frac{(x-45)^2}{12.5}} & -\infty < x < \infty \\ 0 & \text{o.w.} \end{cases}$$

$$\begin{aligned} 1. P(46.7 < x < 48.1) &= P\left(\frac{46.7-\mu}{\sigma} < \frac{x-\mu}{\sigma} < \frac{48.1-\mu}{\sigma}\right) \\ &= P\left(\frac{46.7-45}{2.5} < z < \frac{48.1-45}{2.5}\right) \\ &= P(0.68 < z < 1.24) \\ &= \phi(1.24) - \phi(0.68) = 0.8925 - 0.7517 = 0.1408 \\ 2. P(x > 47.5) &= P\left(\frac{x-\mu}{\sigma} > \frac{47.5-\mu}{\sigma}\right) = P\left(z > \frac{47.5-45}{2.5}\right) \\ &= P(z > 1) = 1 - P(z < 1) \\ &= 1 - \phi(1) = 1 - 0.8413 = 0.1587 \end{aligned}$$

Ex.5: Let $Z \sim SN(0,1)$.

1. Find $P(z < 1.5)$.
2. Find $P(z > 2)$.
3. Find $P(z \leq -1.2)$.
4. Find $P(-0.8 < z < 0.6)$.

Solution:

$$Z \sim N(0,1)$$

$$f(z) = \begin{cases} \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} & -\infty < x < \infty \\ 0 & o.w. \end{cases}$$

1. $P(z < 1.5) = \phi(1.5) = 0.9332$

2. $P(z > 2) = 1 - P(z \leq 2) = 1 - \phi(2) = 1 - 0.9772 = 0.0228$

3. $P(z \leq -1.2) = \phi(-1.2) = 1 - \phi(1.2) = 1 - 0.8849 = 0.1151$

4. $P(-0.8 < z < 0.6) = \phi(0.6) - \phi(-0.8) = \phi(0.6) - [1 - \phi(0.8)]$
 $= 0.7275 - [1 - 0.7881] = 0.5138$

3. The Exponential Distribution

The exponential distribution is one of the important continuous distributions, commonly used in modeling operating times, survival times and waiting times. A random variable X is said to have an exponential distribution if its probability density function (p.d.f.) is given by:

$$f(x, \theta) = \begin{cases} \theta e^{-\theta x} & x > 0 \\ 0 & o.w. \end{cases}$$

Where $\theta > 0$.

1. $f(x, \theta) \geq 0$

2. $\int_{\forall x} f(x, \theta) dx = 1$

$$\begin{aligned} \int_0^{\infty} \theta e^{-\theta x} dx &= - \int_0^{\infty} -\theta e^{-\theta x} dx \\ &= -[e^{-\theta x}]_0^{\infty} = -[e^{-\infty} - e^0] = 1 \end{aligned}$$

3. The **mean** of the r.v. $X \sim \exp(\theta)$ can be obtain as follows

$$\mu_x = E(x) = \frac{1}{\theta}$$

4. The **variance** of the r.v. $X \sim \exp(\theta)$ can be obtain as follows

$$var(x) = \sigma_x^2 = E(x^2) - (E(x))^2 = \frac{1}{\theta^2}$$

5. The Cumulative Distribution Function of X is given by:

$$\begin{aligned} F(x) = P(X \leq x) &= \int_{-\infty}^x f(t) dt = \int_0^x \theta e^{-\theta t} dt \\ &= - \int_0^x -\theta e^{-\theta t} dt = -[e^{-\theta t}]_0^x = -[e^{-\theta x} - e^0] \\ &= -e^{-\theta x} + 1 = 1 - e^{-\theta x} \end{aligned}$$

$$F(x) = \begin{cases} 0 & x < 0 \\ 1 - e^{-\theta x} & 0 \leq x < \infty \end{cases}$$

Remark: Another formula for the exponential distribution is given by:

$$f(x, \theta) = \begin{cases} \frac{1}{\theta} e^{-\frac{x}{\theta}} & x > 0 \\ 0 & o.w. \end{cases}$$

Where $\theta > 0$.

1. $f(x, \theta) \geq 0$

2. $\int_{-\infty}^{\infty} f(x, \theta) dx = 1$

$$\begin{aligned} \int_0^{\infty} \frac{1}{\theta} e^{-\frac{x}{\theta}} dx &= - \int_0^{\infty} -\frac{1}{\theta} e^{-\frac{x}{\theta}} dx \\ &= - \left[e^{-\frac{x}{\theta}} \right]_0^{\infty} = -[e^{-\infty} - e^0] = 1 \end{aligned}$$

3. The **mean** of the r.v. $X \sim \exp(\theta)$ can be obtain as follows

$$\mu_x = E(x) = \theta$$

4. The **variance** of the r.v. $X \sim \exp(\theta)$ can be obtain as follows

$$var(x) = \sigma_x^2 = E(x^2) - (E(x))^2 = \theta^2$$

5. The Cumulative Distribution Function of X is given by:

$$\begin{aligned} F(x) = P(X \leq x) &= \int_{-\infty}^x f(t) dt = \int_0^x \frac{1}{\theta} e^{-\frac{t}{\theta}} dt \\ &= - \int_0^x -\frac{1}{\theta} e^{-\frac{t}{\theta}} dt = - \left[e^{-\frac{t}{\theta}} \right]_0^x = - \left[e^{-\frac{x}{\theta}} - e^0 \right] \\ &= -e^{-\frac{x}{\theta}} + 1 = 1 - e^{-\frac{x}{\theta}} \\ F(x) &= \begin{cases} 0 & x < 0 \\ 1 - e^{-\frac{x}{\theta}} & 0 \leq x < \infty \end{cases} \end{aligned}$$

Ex.6: Suppose that the length of a phone call (in minutes) is an exponential random variable with parameter $\theta = \frac{1}{10}$. If someone arrives immediately ahead of you at a public telephone booth, find the probability that you will have to wait,

1. More than 10 minutes.
2. Between 10 and 20 minutes.

Solution:

$$X \sim \exp\left(\theta = \frac{1}{10}\right)$$

$$f(x, \theta) = \begin{cases} \frac{1}{10} e^{-\frac{x}{10}} & x > 0 \\ 0 & \text{o.w.} \end{cases}$$

$$\begin{aligned} 1. P(x > 10) &= \int_{10}^{\infty} \frac{1}{10} e^{-\frac{x}{10}} dx = - \int_{10}^{\infty} -\frac{1}{10} e^{-\frac{x}{10}} dx \\ &= - \left[e^{-\frac{x}{10}} \right]_{10}^{\infty} = -[e^{-\infty} - e^{-10/10}] = - \left[\frac{1}{e^{\infty}} - e^{-1} \right] \\ &= -[0 - e^{-1}] = e^{-1} = 0.368 \end{aligned}$$

$$\begin{aligned} 2. P(10 < x < 20) &= \int_{10}^{20} \frac{1}{10} e^{-\frac{x}{10}} dx = - \int_{10}^{20} -\frac{1}{10} e^{-\frac{x}{10}} dx \\ &= - \left[e^{-\frac{x}{10}} \right]_{10}^{20} = -[e^{-20/10} - e^{-10/10}] = -[e^{-2} - e^{-1}] \\ &= e^{-1} - e^{-2} = 0.233 \end{aligned}$$

Ex.7: Let $X \sim \exp(2)$.

1. Write the p.d.f. of X .
2. Find the mean and the variance of X .
3. Find the cumulative distribution function of X .
4. $P(x > 2), P(3 < x < 5), P(|x| < 3)$.
5. $P(2 < x \leq 4)$ using c.d.f.

Solution:

$$1. X \sim \exp(2)$$

$$f(x, 2) = \begin{cases} 2 e^{-2x} & x > 0 \\ 0 & \text{o.w.} \end{cases}$$

$$2. \mu_x = E(x) = \frac{1}{\theta} = \frac{1}{2}, \quad \text{var}(x) = \sigma_x^2 = \frac{1}{\theta^2} = \frac{1}{4}$$

3.

$$F(x) = \begin{cases} 0 & x < 0 \\ 1 - e^{-2x} & 0 \leq x < \infty \end{cases}$$

4.

$$\begin{aligned} \Rightarrow P(x > 2) &= \int_2^{\infty} 2 e^{-2x} dx = - \int_2^{\infty} -2 e^{-2x} dx \\ &= -[e^{-2x}]_2^{\infty} = -[e^{-\infty} - e^{-4}] = e^{-4} = 0.0183 \end{aligned}$$

$$\begin{aligned} \Rightarrow P(3 < x < 5) &= \int_3^5 2 e^{-2x} dx = - \int_3^5 -2 e^{-2x} dx \\ &= -[e^{-2x}]_3^5 = -[e^{-10} - e^{-6}] = e^{-6} - e^{-10} = 0.00243 \end{aligned}$$

$$\begin{aligned} \Rightarrow P(|x| < 3) &= P(-3 < x < 3) = P(-3 < x < 0) + P(0 < x < 3) \\ &= 0 + P(0 < x < 3) = P(0 < x < 3) \end{aligned}$$

$$\begin{aligned} &= \int_0^3 2 e^{-2x} dx = - \int_0^3 -2 e^{-2x} dx = -[e^{-2x}]_0^3 = -[e^{-3} - e^0] \\ &= 1 - e^{-3} = 0.95 \end{aligned}$$

$$5. P(2 < x \leq 4) = F(4) - F(2)$$

$$= (1 - e^{-8}) - (1 - e^{-4}) = e^{-4} - e^{-8} = 0.01797$$

Ex.8: Let the r.v. X denote the operating times (in hours) of a machine, if $X \sim \text{exp}(0.04)$. Answer each of the following:

1. Write the p.d.f. of X .
2. Write the c.d.f. of X .
3. Find the mean of machine operating times.
4. Find the probability that the machine will work at least 20 hours.

Solution:

1. $X \sim \exp(0.04)$

$$f(x, 0.04) = \begin{cases} 0.04 e^{-0.04x} & x > 0 \\ 0 & o.w. \end{cases}$$

2.

$$F(x) = \begin{cases} 0 & x < 0 \\ 1 - e^{-0.04x} & 0 \leq x < \infty \end{cases}$$

3. $\mu_x = \frac{1}{\theta} = \frac{1}{0.04} = 25 \text{ hours}$

4. $P(x \geq 20) = \int_{20}^{\infty} 0.04 e^{-0.04x} dx = - \int_{20}^{\infty} -0.04 e^{-0.04x} dx$
 $= -[e^{-0.04x}]_{20}^{\infty} = -[e^{-\infty} - e^{-0.8}] = e^{-0.8} = 0.449$

Ex.9: Let the r.v. X denote the operating times (in hours) of a machine, suppose the average operating time of the machine is 25 hours. Answer each of the following:

1. Write the p.d.f. of X .
2. Write the c.d.f. of X .
3. Find the mean of machine operating times.
4. Find the probability that the machine will work at least 20 hours.

Solution:

1. $X \sim \exp(25)$

$$f(x, 0.04) = \begin{cases} \frac{1}{25} e^{-x/25} & x > 0 \\ 0 & o.w. \end{cases}$$

2.

$$F(x) = \begin{cases} 0 & x < 0 \\ 1 - e^{-x/25} & 0 \leq x < \infty \end{cases}$$

3. $\mu_x = \theta = 25 \text{ hours}$

4. $P(x \geq 20) = \int_{20}^{\infty} \frac{1}{25} e^{-x/25} dx = - \int_{20}^{\infty} -\frac{1}{25} e^{-x/25} dx$
 $= -[e^{-x/25}]_{20}^{\infty} = -[e^{-\infty} - e^{-0.8}] = e^{-0.8} = 0.449$

4. Gamma Distribution

A random variable X is said to have a gamma distribution with parameters α and β if its probability density function (p.d.f.) is given by:

$$f(x, \alpha, \beta) = \begin{cases} \frac{\beta^\alpha}{\Gamma\alpha} x^{\alpha-1} e^{-\beta x} & x > 0 \\ 0 & \text{o.w.} \end{cases}$$

$$\Gamma\alpha = \int_0^{\infty} x^{\alpha-1} e^{-x} dx = (\alpha - 1)! \quad \text{Gamma Function}$$

Where the parameters $\alpha, \beta > 0$.

1. $f(x, \alpha, \beta) \geq 0$
2. $\int_{\forall x} f(x, \alpha, \beta) dx = 1$

$$\int_0^{\infty} \frac{\beta^\alpha}{\Gamma\alpha} x^{\alpha-1} e^{-\beta x} dx = \frac{\beta^\alpha}{\Gamma\alpha} \int_0^{\infty} x^{\alpha-1} e^{-\beta x} dx$$

$$\text{Let } y = \beta x \Rightarrow x = \frac{y}{\beta} \Rightarrow dx = \frac{1}{\beta} dy$$

Substitute

$$\frac{\beta^\alpha}{\Gamma\alpha} \int_0^{\infty} \left(\frac{y}{\beta}\right)^{\alpha-1} e^{-y} \frac{1}{\beta} dy$$

$$\frac{\beta^\alpha}{\Gamma\alpha} * \frac{1}{\beta^\alpha} \int_0^{\infty} y^{\alpha-1} e^{-y} dy$$

$$\frac{1}{\Gamma\alpha} \int_0^{\infty} y^{\alpha-1} e^{-y} dy$$

$$\Gamma\alpha = \int_0^{\infty} y^{\alpha-1} e^{-y} dy$$

$$\frac{1}{\Gamma\alpha} * \Gamma\alpha = 1$$

3. The **mean** of the r.v. $X \sim \text{Gamm}(\alpha, \beta)$ can be obtain as follows

$$\mu_x = E(x) = \frac{\alpha}{\beta}$$

4. The **variance** of the r.v. $X \sim \text{Gamm}(\alpha, \beta)$ can be obtain as follows

$$\text{var}(x) = \sigma_x^2 = E(x^2) - (E(x))^2 = \frac{\alpha}{\beta^2}$$

5. The Cumulative Distribution Function of X is given by:

$$F(x) = P(X \leq x) = \int_{-\infty}^x f(t) dt = \int_0^x \frac{\beta^\alpha}{\Gamma\alpha} t^{\alpha-1} e^{-\beta t} dt$$

Remark 1: Another formula for the Gamma distribution is given by:

$$f(x, \alpha, \beta) = \begin{cases} \frac{1}{\Gamma\alpha \beta^\alpha} x^{\alpha-1} e^{-x/\beta} & x > 0 \\ 0 & \text{o.w.} \end{cases}$$

$$\Gamma\alpha = \int_0^{\infty} x^{\alpha-1} e^{-x} dx = (\alpha - 1)! \quad \text{Gamma Function}$$

Where the parameters $\alpha, \beta > 0$.

$$1. f(x, \alpha, \beta) \geq 0$$

$$2. \int_{\forall x} f(x, \alpha, \beta) dx = 1$$

$$\int_0^{\infty} \frac{1}{\Gamma\alpha \beta^\alpha} x^{\alpha-1} e^{-x/\beta} dx = \frac{1}{\Gamma\alpha \beta^\alpha} \int_0^{\infty} x^{\alpha-1} e^{-x/\beta} dx$$

$$\text{Let } y = \frac{x}{\beta} \Rightarrow x = \beta y \Rightarrow dx = \beta dy$$

Substitute

$$\frac{1}{\Gamma\alpha \beta^\alpha} \int_0^{\infty} (\beta y)^{\alpha-1} e^{-y} \beta dy$$

$$\frac{\beta^\alpha}{\Gamma\alpha \beta^\alpha} \int_0^{\infty} y^{\alpha-1} e^{-y} dy$$

$$\frac{1}{\Gamma\alpha} \int_0^{\infty} y^{\alpha-1} e^{-y} dy$$

$$\Gamma\alpha = \int_0^{\infty} y^{\alpha-1} e^{-y} dy$$

$$\frac{1}{\Gamma\alpha} * \Gamma\alpha = 1$$

3. The **mean** of the r.v. $X \sim \text{Gamm}(\alpha, \beta)$ can be obtain as follows

$$\mu_x = E(x) = \alpha\beta$$

4. The **variance** of the r.v. $X \sim \text{Gamm}(\alpha, \beta)$ can be obtain as follows

$$\text{var}(x) = \sigma_x^2 = E(x^2) - (E(x))^2 = \alpha\beta^2$$

5. The Cumulative Distribution Function of X is given by:

$$F(x) = P(X \leq x) = \int_{-\infty}^x f(t) dt = \int_0^x \frac{1}{\Gamma\alpha \beta^\alpha} t^{\alpha-1} e^{-t/\beta} dt$$

Remark 2: If $\alpha = 1$, X is said to have an exponential distribution with parameter β , $X \sim \text{exp}(\beta)$

$$f(x, \alpha, \beta) = \frac{\beta^\alpha}{\Gamma\alpha} x^{\alpha-1} e^{-\beta x}$$

$$f(x, \beta) = \begin{cases} \beta e^{-\beta x} & x > 0 \\ 0 & \text{o.w.} \end{cases}$$

Remark 3: If $\alpha = \frac{n}{2}$ and $\beta = \frac{1}{2}$, X is said to have an chi-square distribution with parameter n , $X \sim \chi^2(n)$

$$f(x, \alpha, \beta) = \frac{\left(\frac{1}{2}\right)^{n/2}}{\Gamma \frac{n}{2}} x^{\frac{n}{2}-1} e^{-\frac{1}{2}x}$$

$$f(x, n) = \begin{cases} \frac{1}{2^{n/2} \Gamma \frac{n}{2}} x^{\frac{n}{2}-1} e^{-\frac{1}{2}x} & x > 0 \\ 0 & o.w. \end{cases}$$

Ex.10: Let $X \sim \text{Gamm}(2,4)$

1. Write the p.d.f. of X .
2. Find $P(x < 3)$.

Solution:

$$1. f(x, \alpha, \beta) = \begin{cases} \frac{\beta^\alpha}{\Gamma \alpha} x^{\alpha-1} e^{-\beta x} & x > 0 \\ 0 & o.w. \end{cases}$$

$$f(x, \alpha, \beta) = \frac{4^2}{\Gamma 2} x^{2-1} e^{-4x}$$

$$\Gamma 2 = (2 - 1)! = 1$$

$$f(x, \alpha, \beta) = \begin{cases} 16 x e^{-4x} & x > 0 \\ 0 & o.w. \end{cases}$$

$$2. P(x < 3) = \int_0^3 16 x e^{-4x} dx = 16 \int_0^3 x e^{-4x} dx$$

$$\int u dv = uv - \int v du$$

$$u = x \quad , \quad du = dx$$

$$dv = e^{-4x} \quad , \quad v = \frac{-1}{4} e^{-4x}$$

$$\begin{aligned} \int_0^3 x e^{-4x} dx &= \left[x \frac{-1}{4} e^{-4x} \right]_0^3 + \int_0^3 \frac{-1}{4} e^{-4x} dx \\ &= \left[\frac{-3}{4} e^{-12} - 0 \right] - \frac{1}{4} * \frac{1}{4} \int_0^3 (-4) e^{-4x} dx \\ &= \frac{-3}{4} e^{-12} - \frac{1}{16} [e^{-4x}]_0^3 \\ &= \frac{-3}{4} e^{-12} - \frac{1}{16} [e^{-12} - e^0] \end{aligned}$$

$$= \frac{-3}{4} e^{-12} - \frac{1}{16} [e^{-12} - e^0] = 0.062$$

$$P(x < 3) = 16 \int_0^3 x e^{-4x} dx = 16 * 0.062 = 0.992$$

Ex.11: Let X be a random variable with the following p.d.f.

$$f(x, \alpha, \beta) = \begin{cases} A x^2 e^{-5x} & x > 0 \\ 0 & o.w. \end{cases}$$

Find the value of A , what is the name of the distribution.

Solution:

$$f(x, \alpha, \beta) = \frac{\beta^\alpha}{\Gamma\alpha} x^{\alpha-1} e^{-\beta x}$$

$$f(x, \alpha, \beta) = A x^2 e^{-5x}$$

$$\beta = 5, \alpha - 1 = 2 \Rightarrow \alpha = 3$$

$$A = \frac{\beta^\alpha}{\Gamma\alpha} = \frac{5^3}{\Gamma 3} = \frac{125}{2!} = \frac{125}{2}$$

$$f(x, \alpha, \beta) = \begin{cases} \frac{125}{2} x^2 e^{-5x} & x > 0 \\ 0 & o.w. \end{cases}$$

$\therefore X \sim \text{Gamm}(3,5)$

Ex.12: The lifetime (in hours) of a battery follows a Gamma distribution with $\alpha = 3$ and $\beta = 5$ (*scale parameter*)

1. Write the probability density function of X .
2. Find the mean and the variance of X .

Solution:

$$1. f(x, \alpha, \beta) = \frac{1}{\Gamma 3 5^3} x^{3-1} e^{-x/5}$$

$$\Gamma 3 = (3 - 1)! = 2! = 2$$

$$f(x, \alpha, \beta) = \begin{cases} \frac{1}{250} x^2 e^{-x/5} & x > 0 \\ 0 & \text{o.w.} \end{cases}$$

$$2. \mu_x = \alpha\beta = 3 * 5 = 15, \quad \text{var}(x) = \alpha\beta^2 = 3 * 5^2 = 75$$

Ex.13: The time (in hours) until a certain machine fails follows a Gamma distribution with $\alpha = 4$ and rate parameter $\beta = 2$ per hour.

1. Write the probability density function of X .
2. Find the mean and the variance of X .

Solution:

$$1. f(x, \alpha, \beta) = \frac{2^4}{\Gamma 4} x^{4-1} e^{-2x}$$

$$\Gamma 4 = (4 - 1)! = 3! = 6$$

$$f(x, \alpha, \beta) = \begin{cases} \frac{16}{6} x^3 e^{-2x} & x > 0 \\ 0 & \text{o.w.} \end{cases}$$

$$2. \mu_x = \frac{\alpha}{\beta} = \frac{4}{2} = 2, \quad \text{var}(x) = \frac{\alpha}{\beta^2} = \frac{4}{2^2} = 1$$

5. Beta Distribution

A random variable X is said to have a Beta distribution with parameters a and b if its probability density function (p.d.f.) is given by:

$$f(x, a, b) = \begin{cases} \frac{1}{B(a, b)} x^{a-1} (1-x)^{b-1} & 0 < x < 1 \\ 0 & o.w. \end{cases}$$

$$B(a, b) = \int_0^1 x^{a-1} (1-x)^{b-1} dx = \frac{\Gamma a \Gamma b}{\Gamma a + b} \quad \text{Beta Function}$$

Where the parameters $a, b > 0$.

1. $f(x, a, b) \geq 0$
2. $\int_{\forall x} f(x, a, b) dx = 1$

$$\int_0^1 \frac{1}{B(a, b)} x^{a-1} (1-x)^{b-1} dx$$

$$\frac{1}{B(a, b)} \int_0^1 x^{a-1} (1-x)^{b-1} dx$$

$$\frac{1}{B(a, b)} B(a, b) = 1$$

3. The **mean** of the r.v. $X \sim \text{Beta}(a, b)$ can be obtain as follows

$$\mu_x = E(x) = \frac{a}{a+b}$$

4. The **variance** of the r.v. $X \sim \text{Beta}(a, b)$ can be obtain as follows

$$\text{var}(x) = \sigma_x^2 = E(x^2) - (E(x))^2 = \frac{ab}{(a+b)^2 (a+b+1)}$$

5. The Cumulative Distribution Function of X is given by:

$$F(x) = P(X \leq x) = \int_{-\infty}^x f(t) dt = \int_0^x \frac{1}{B(a, b)} t^{a-1} (1-t)^{b-1} dt$$

Ex.14: Let $X \sim \text{Beta}(2,3)$

1. Write the p.d.f. of X .
2. Find the c.d.f. of X .
3. Find the mean and the variance of X .
4. Find $P\left(\frac{1}{2} < x < 1\right), P(-1 < x < 1), P(x > 0)$.

Solution:

$$1. f(x, a, b) = \frac{1}{B(a,b)} x^{a-1} (1-x)^{b-1}$$

$$f(x, a, b) = \frac{1}{B(2,3)} x^{2-1} (1-x)^{3-1}$$

$$B(2,3) = \frac{\Gamma 2 \Gamma 3}{\Gamma 2 + 3} = \frac{1 * 2}{24} = \frac{1}{12}$$

$$\Gamma 2 = (2 - 1)! = 1! = 1$$

$$\Gamma 3 = (3 - 1)! = 2! = 2$$

$$\Gamma 5 = (5 - 1)! = 4! = 24$$

$$f(x, a, b) = \begin{cases} \frac{1}{1/12} x (1-x)^2 & 0 < x < 1 \\ 0 & \text{o.w.} \end{cases}$$

$$f(x, a, b) = \begin{cases} 12 x (1-x)^2 & 0 < x < 1 \\ 0 & \text{o.w.} \end{cases}$$

2.

$$F(x) = \int_{-\infty}^x f(t) dt = \int_0^x 12 t (1-t)^2 dt$$

$$= 12 \int_0^x t (1 - 2t + t^2) = 12 \int_0^x t - 2t^2 + t^3$$

$$= 12 \left[\frac{t^2}{2} - 2 \frac{t^3}{3} + \frac{t^4}{4} \right]_0^x = 12 \left[\left(\frac{x^2}{2} - 2 \frac{x^3}{3} + \frac{x^4}{4} \right) - 0 \right]$$

$$F(x) = \begin{cases} 0 & x < 0 \\ 6x^2 - 8x^3 + 3x^4 & 0 \leq x < 1 \\ 1 & x \geq 1 \end{cases}$$

$$3. \mu_x = \frac{a}{a+b} = \frac{2}{2+3} = \frac{2}{5}$$

$$\text{var}(x) = \frac{ab}{(a+b)^2 (a+b+1)} = \frac{2 * 3}{(2+3)^2 (2+3+1)} = \frac{6}{150}$$

$$4. P\left(\frac{1}{2} < x < 1\right) = F(1) - F\left(\frac{1}{2}\right) = 1 - 0.6875 = 0.3125$$

$$F(1) = 1$$

$$F\left(\frac{1}{2}\right) = 6\left(\frac{1}{2}\right)^2 - 8\left(\frac{1}{2}\right)^3 + 3\left(\frac{1}{2}\right)^4 = 0.6875$$

$$P(-1 < x < 1) = P(-1 < x < 0) + P(0 < x < 1) = P(0 < x < 1) \\ = F(1) - F(0) = 1 - 0 = 1$$

$$P(x > 0) = 1 - P(x \leq 0) = 1 - F(0) = 1 - 0 = 1$$

Ex.15: Let X be a random variable with the following p.d.f.

$$f(x, \alpha, \beta) = \begin{cases} A x^3 (1-x)^2 & 0 < x < 1 \\ 0 & \text{o. w.} \end{cases}$$

1. Find the value of A .
2. What is the name of the distribution?
3. Find the c.d.f. of X .
4. Find $P\left(0 < x < \frac{1}{2}\right), P(x < 1), P\left(\frac{1}{2} < x < 3\right)$.

Solution:

$$1. f(x, a, b) = \frac{1}{B(a,b)} x^{a-1} (1-x)^{b-1}$$

$$f(x, a, b) = A x^3 (1-x)^2$$

$$a - 1 = 3 \Rightarrow a = 4, b - 1 = 2 \Rightarrow b = 3$$

$$A = \frac{1}{B(a,b)} = \frac{1}{\frac{\Gamma a \Gamma b}{\Gamma a + b}} = \frac{\Gamma a + b}{\Gamma a \Gamma b} = \frac{\Gamma 4 + 3}{\Gamma 4 \Gamma 3} = \frac{720}{6 * 2} = \frac{720}{12} = 60$$

$$\Gamma 4 + 3 = \Gamma 7 = (7 - 1)! = 6! = 720$$

$$\Gamma 4 = (4 - 1)! = 3! = 6$$

$$\Gamma 3 = (3 - 1)! = 2! = 2$$

$$f(x, a, b) = \begin{cases} 60 x^3 (1-x)^2 & 0 < x < 1 \\ 0 & \text{o.w.} \end{cases}$$

2. $X \sim \text{Beta}(4,3)$

$$\begin{aligned} 3. F(x) &= \int_{-\infty}^x f(t) dt = \int_0^x 60 t^3 (1-t)^2 dt \\ &= 60 \int_0^x t^3 (1-2t+t^2) dt = 60 \int_0^x t^3 - 2t^4 + t^5 dt \\ &= 60 \left[\frac{t^4}{4} - 2\frac{t^5}{5} + \frac{t^6}{6} \right]_0^x = 60 \left[\left(\frac{x^4}{4} - 2\frac{x^5}{5} + \frac{x^6}{6} \right) - 0 \right] \end{aligned}$$

$$F(x) = \begin{cases} 0 & x < 0 \\ 15x^4 - 24x^5 + 10x^6 & 0 \leq x < 1 \\ 1 & x \geq 1 \end{cases}$$

$$4. P\left(0 < x < \frac{1}{2}\right) = F\left(\frac{1}{2}\right) - F(0) = 0.34375 - 0 = 0.34375$$

$$F\left(\frac{1}{2}\right) = 15\left(\frac{1}{2}\right)^4 - 24\left(\frac{1}{2}\right)^5 + 10\left(\frac{1}{2}\right)^6 = 0.34375$$

$$F(0) = 15(0)^4 - 24(0)^5 + 10(0)^6 = 0$$

$$P(x < 1) = F(1) = 15(1)^4 - 24(1)^5 + 10(1)^6 = 1$$

$$\begin{aligned} P\left(\frac{1}{2} < x < 3\right) &= P\left(\frac{1}{2} < x < 1\right) + P(1 < x < 3) = P\left(\frac{1}{2} < x < 1\right) \\ &= F(1) - F\left(\frac{1}{2}\right) = 1 - 0.34375 = 0.65625 \end{aligned}$$

6. Weibull Distribution

A random variable X is said to have a Weibull distribution with parameters α and λ if its probability density function (p.d.f.) is given by:

$$f(x, \alpha, \lambda) = \begin{cases} \frac{\alpha}{\lambda} \left(\frac{x}{\lambda}\right)^{\alpha-1} e^{-\left(\frac{x}{\lambda}\right)^\alpha} & x \geq 0 \\ 0 & o.w. \end{cases}$$

Where the parameters $\alpha, \lambda > 0$.

1. $f(x, \alpha, \lambda) \geq 0$
2. $\int_{-\infty}^{\infty} f(x, \alpha, \lambda) dx = 1$
3. The **mean** of the r.v. $X \sim Wei(\alpha, \lambda)$ can be obtain as follows

$$\mu_x = E(x) = \lambda \Gamma\left(1 + \frac{1}{\alpha}\right)$$

4. The **variance** of the r.v. $X \sim Wei(\alpha, \lambda)$ can be obtain as follows

$$var(x) = \sigma_x^2 = E(x^2) - (E(x))^2 = \lambda^2 \left[\Gamma\left(1 + \frac{2}{\alpha}\right) - \left[\Gamma\left(1 + \frac{1}{\alpha}\right) \right]^2 \right]$$

5. The Cumulative Distribution Function of X is given by:

$$F(x) = P(X \leq x) = \int_{-\infty}^x f(t) dt = \int_0^x \frac{\alpha}{\lambda} \left(\frac{t}{\lambda}\right)^{\alpha-1} e^{-\left(\frac{t}{\lambda}\right)^\alpha} dt$$

$$F(x) = \begin{cases} 0 & x < 0 \\ 1 - e^{-\left(\frac{x}{\lambda}\right)^\alpha} & x \geq 0 \end{cases}$$

Ex.16: Suppose the life (in hours) of a certain type of light bulb follows a Weibull distribution with parameters $\alpha = 2$ and $\lambda = 1000$ hours.

($\Gamma 1.5 = 0.8862$,

1. Write the p.d.f. of X .
2. Find the c.d.f. of X .
3. Find the mean and the variance of X .
4. Find the probability that a bulb lasts less than 800 hours.

Solution:

$$1. f(x, \alpha, \lambda) = \begin{cases} \frac{2}{1000} \left(\frac{x}{1000}\right)^{2-1} e^{-\left(\frac{x}{1000}\right)^2} & x \geq 0 \\ 0 & o.w. \end{cases}$$

$$2. F(x) = \begin{cases} 0 & x < 0 \\ 1 - e^{-\left(\frac{x}{1000}\right)^2} & x \geq 0 \end{cases}$$

$$3. \mu_x = \lambda \Gamma\left(1 + \frac{1}{\alpha}\right) = 1000 \Gamma\left(1 + \frac{1}{2}\right) = 1000 \Gamma\frac{3}{2} \\ = 1000 * \Gamma 1.5 = 1000 * 0.8862 = 886.2 \text{ hours}$$

$$\begin{aligned} var(x) &= \lambda^2 \left[\Gamma\left(1 + \frac{2}{\alpha}\right) - \left[\Gamma\left(1 + \frac{1}{\alpha}\right) \right]^2 \right] \\ &= 1000^2 \left[\Gamma\left(1 + \frac{2}{2}\right) - \left[\Gamma\left(1 + \frac{1}{2}\right) \right]^2 \right] \\ &= 1000^2 [\Gamma(2) - [\Gamma(1.5)]^2] = 1000000(1 - (0.8862)^2) \\ &= 214.700 \text{ hours} \end{aligned}$$

$$4. P(x < 800) = F(800) = 1 - e^{-\left(\frac{800}{1000}\right)^2} = 1 - e^{-0.64} \\ = 1 - 0.527 = 0.473$$

7. Pareto Distribution

A random variable X is said to have a Pareto distribution with parameter θ if its probability density function (p.d.f.) is given by:

$$f(x, \theta) = \begin{cases} \frac{\theta}{x^{\theta+1}} & x \geq 1 \\ 0 & \text{o.w.} \end{cases}$$

Where the parameters $\theta > 0$.

1. $f(x, \theta) \geq 0$
2. $\int_{\forall x} f(x, \theta) dx = 1$
3. The **mean** of the r.v. $X \sim Par(\theta)$ can be obtain as follows

$$\mu_x = E(x) = \frac{\theta}{\theta - 1}$$

4. The **variance** of the r.v. $X \sim Par(\theta)$ can be obtain as follows

$$var(x) = \sigma_x^2 = E(x^2) - (E(x))^2 = \frac{\theta}{(\theta - 1)^2 (\theta - 2)}$$

5. The Cumulative Distribution Function of X is given by:

$$F(x) = P(X \leq x) = \int_{-\infty}^x f(t) dt = \int_1^x \frac{\theta}{t^{\theta+1}} dt$$

$$F(x) = \begin{cases} 0 & x < 1 \\ 1 - \left(\frac{1}{x}\right)^\theta & x \geq 1 \end{cases}$$

Ex.17: Let X be a random variable that follows Pareto distribution with parameter $\theta = 3$.

1. Write the p.d.f. of X .
2. Find the c.d.f. of X .
3. Find the mean and the variance of X .
4. Find $P(x > 2)$.

Solution:

$$1. f(x, \theta) = \begin{cases} \frac{3}{x^4} & x \geq 1 \\ 0 & \text{o.w.} \end{cases}$$

$$2. F(x) = \begin{cases} 0 & x < 1 \\ 1 - \left(\frac{1}{x}\right)^3 & x \geq 1 \end{cases}$$

$$3. \mu_x = \frac{\theta}{\theta-1} = \frac{3}{3-1} = \frac{3}{2}$$

$$\text{var}(x) = \frac{\theta}{(\theta-1)^2(\theta-2)} = \frac{3}{(3-1)^2(3-2)} = \frac{3}{4}$$

$$4. P(x > 2) = 1 - P(x \leq 2) = 1 - F(2) = 1 - 1 + \left(\frac{1}{2}\right)^3 = \frac{1}{8}$$

8. Rayleigh Distribution

A random variable X is said to have a Rayleigh distribution with parameter λ if its probability density function (p.d.f.) is given by:

$$f(x, \lambda) = \begin{cases} \frac{x}{\lambda^2} e^{-\frac{x^2}{2\lambda^2}} & x \geq 0 \\ 0 & o.w. \end{cases}$$

Where the parameters $\lambda > 0$.

1. $f(x, \lambda) \geq 0$
2. $\int_{-\infty}^{\infty} f(x, \lambda) dx = 1$
3. The **mean** of the r.v. $X \sim Ray(\lambda)$ can be obtain as follows

$$\mu_x = E(x) = \sqrt{\frac{\pi}{2}} \lambda$$

4. The **variance** of the r.v. $X \sim Par(\theta)$ can be obtain as follows

$$var(x) = \sigma_x^2 = E(x^2) - (E(x))^2 = \frac{4 - \pi}{2} \lambda^2$$

5. The Cumulative Distribution Function of X is given by:

$$F(x) = P(X \leq x) = \int_{-\infty}^x f(t) dt = \int_0^x \frac{t}{\lambda^2} e^{-\frac{t^2}{2\lambda^2}} dt$$

$$F(x) = \begin{cases} 0 & x < 0 \\ 1 - e^{-\frac{x^2}{2\lambda^2}} & x \geq 0 \end{cases}$$

Ex.18: A machine measures vibration, and the vibration follows a Rayleigh distribution with $\lambda = 2$.

1. Write the p.d.f. of X .
2. Find the c.d.f. of X .
3. Find the mean and the variance of X .
4. What is the probability that the vibration is less than 2.

Solution:

$$1. f(x, \lambda) = \begin{cases} \frac{x}{4} e^{-\frac{x^2}{8}} & x \geq 0 \\ 0 & o.w. \end{cases}$$

$$2. F(x) = \begin{cases} 0 & x < 0 \\ 1 - e^{-\frac{x^2}{8}} & x \geq 0 \end{cases}$$

$$3. \mu_x = \lambda \sqrt{\frac{\pi}{2}} = 2 * \sqrt{\frac{3.14}{2}} = 2 * 1.2533 = 2.50$$

$$var(x) = \frac{4 - \pi}{2} \lambda^2 = \frac{4 - 3.14}{2} * 2^2 = 1.716$$

$$4. P(x < 2) = F(2) = 1 - e^{-\frac{2^2}{8}} = 1 - e^{-0.5} = 1 - 0.6065 = 0.3935$$